Model-based clustering

Gaussian mixture models

June 15th, 2023

Previously...

- We explored the use of K-means and hierarchical clustering for clustering
- These methods yield **hard** assignments, strictly assigning observations to only one cluster
- What about **soft** assignments? Allow for some **uncertainty** in the clustering results
- Welcome to the wonderful world of **mixture models**



Previously in kernel density estimation...

$$ext{Kernel density estimate:} \ \hat{f}\left(x
ight) = rac{1}{n}\sum_{i=1}^{n}rac{1}{h}K_{h}(x-x_{i})$$

• We have to use every observation when estimating the density for new points



• Instead we can make **assumptions** to "simplify" the problem

Mixture models

We assume the distribution f(x) is a **mixture** of *K* component distributions:

$$f(x) = \sum_{k=1}^K \pi_k f_k(x)$$

• $\pi_k =$ mixing proportions (or weights), where $\pi_k > 0$, and $\sum_k \pi_k = 1$

This is a **data generating process**, meaning to generate a new point:

1. pick a distribution / component among our K options, by introducing a new variable:

 $\cdot z \sim ext{Multinomial}(\pi_1, \pi_2, \dots, \pi_k)$, i.e. categorical variable saying which group the new point is from

2. generate an observation with that distribution / component, i.e. $x|z\sim f_z$

So what do we use for each f_k ?

Gaussian mixture models (GMMs)

Assume a **parametric mixture model**, with **parameters** θ_k for the kth component

$$f(x) = \sum_{k=1}^K \pi_k f_k(x; heta_k)$$

Assume each component is Gaussian / Normal where for 1D case:

$$f_k(x; heta_k)=N(x;\mu_k,\sigma_k^2)=rac{1}{\sqrt{2\pi\sigma_k^2}}{
m exp}\Big(-rac{(x-\mu_k)^2}{2\sigma_k^2}\Big)$$

We need to estimate each $\pi_1, \ldots, \pi_k, \mu_1, \ldots, \mu_k, \sigma_1, \ldots, \sigma_k!$

Let's pretend we only have one component...

If we have *n* observations from a single Normal distribution, we estimate the distribution parameters using the **likelihood function**, the probability / density of observing the data given the parameters

$$\mathcal{L}(\mu,\sigma|x_1,\ldots,x_n)=f(x_1,\ldots,x_n|\mu,\sigma)=\prod_i^nrac{1}{\sqrt{2\pi\sigma^2}}\mathrm{exp}~-rac{(x_i-\mu)^2}{2\sigma^2}$$

We can compute the **maximum likelihood estimates (MLEs)** for μ and σ

You already know these values!

•
$$\hat{\mu}_{MLE} = rac{1}{n}\sum_{i}^{n}x_{i}$$
, sample mean

•
$$\hat{\sigma}_{MLE} = \sqrt{rac{1}{n}\sum_{i}^{n}(x_i-\mu)^2}$$
, sample standard deviation (plug in $\hat{\mu}_{MLE}$)

The problem with more than one component...

- We don't know which component an observation belongs to
- **IF WE DID KNOW**, then we could compute each component's MLEs as before
- But we don't know because *z* is a **latent variable**! So what about its distribution given the data?

$$egin{aligned} P(z_i = k | x_i) &= rac{P(x_i | z_i = k) P(z_i = k)}{P(x_i)} \ &= rac{\pi_k N\left(\mu_k, \sigma_k^2
ight)}{\sum_{k=1}^K \pi_k N\left(\mu_k, \sigma_k
ight)} \end{aligned}$$

- But we do NOT know these parameters!
- This leads to a very useful algorithm in statistics...



Expectation-maximization (EM) algorithm

We alternate between the following:

- *pretending* to know the probability each observation belongs to each group, to estimate the parameters of the components
- *pretending* to know the parameters of the components, to estimate the probability each observation belong to each group

Where have you seen this before? K-means algorithm!

- 1. Start with initial guesses about $\pi_1, \ldots, \pi_k, \mu_1, \ldots, \mu_k, \sigma_1, \ldots, \sigma_k$
- 2. Repeat until nothing changes:
- Expectation step: calculate \hat{z}_{ik} = expected membership of observation i in cluster k
- Maximization step: update parameter estimates with weighted MLE using \hat{z}_{ik}

How does this relate back to clustering?

From the EM algorithm: \hat{z}_{ik} is a **soft membership** of observation *i* in cluster *k*

- you can assign observation i to a cluster with the largest \hat{z}_{ik}
- measure cluster assignment $\mathbf{uncertainty} = 1 \max_k \hat{z}_{ik}$

Our parameters determine the type of clusters

In 1D we only have two options:

1. each cluster is assumed to have equal variance (spread): $\sigma_1^2 = \sigma_2^2 = \cdots = \sigma_k^2$

2. each cluster **is allowed to have a different variance**

But that is only 1D... what happens in multiple dimensions?

Multivariate GMMs

$$f(x) = \sum_{k=1}^K \pi_k f_k(x; heta_k)$$
 where $f_k(x; heta_k) \sim N(oldsymbol{\mu}_k, oldsymbol{\Sigma}_k)$

Each component is a **multivariate normal distribution**:

- $\boldsymbol{\mu}_k$ is a *vector* of means in p dimensions
- $\mathbf{\Sigma}_k$ is the p imes p covariance matrix describes the joint variability between pairs of variables

$$\sum = \begin{bmatrix} \sigma_1^2 & \sigma_{1,2} & \cdots & \sigma_{1,p} \\ \sigma_{2,1} & \sigma_2^2 & \cdots & \sigma_{2,p} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{p,1} & \sigma_{p,2}^2 & \cdots & \sigma_p^2 \end{bmatrix}$$

Covariance constraints

$$\sum = \begin{bmatrix} \sigma_1^2 & \sigma_{1,2} & \cdots & \sigma_{1,p} \\ \sigma_{2,1} & \sigma_2^2 & \cdots & \sigma_{2,p} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{p,1} & \sigma_{p,2}^2 & \cdots & \sigma_p^2 \end{bmatrix}$$

As we increase the number of dimensions, model fitting and estimation becomes increasingly difficult

We can use **constraints** on multiple aspects of the k covariance matrices:

- volume: size of the clusters, i.e., number of observations,
- shape: direction of variance, i.e. which variables display more variance
- orientation: aligned with axes (low covariance) versus tilted (due to relationships between variables)



- Control volume, shape, orientation
- E means equal and V means variable (VVV is the most flexible, but has the most parameters)
- Two II is **spherical**, one I is **diagonal**, and the remaining are **general**

So many options! How do we know what to do?



Bayesian information criterion (BIC)

This is a statistical model

$$f(x) = \sum_{k=1}^K \pi_k f_k(x; heta_k) \ ext{where} \ f_k(x; heta_k) \sim N(oldsymbol{\mu}_k, oldsymbol{\Sigma}_k)$$

Meaning we can use a **model selection** procedure for determining which best characterizes the data Specifically - we will use a **penalized likelihood** measure

$$BIC = 2\log \mathcal{L} - m\log n$$

- $\log \mathcal{L}$ is the log-likelihood of the considered model
- with m parameters (VVV has the most parameters) and n observations
- **penalizes** large models with **many clusters without constraints**
- we can use BIC to choose the covariance constraints AND number of clusters K!

The above BIC is really the -BIC of what you typically see, this sign flip is just for ease

Mixture model for NBA players... New dataset!

Created dataset of NBA player statistics per 100 possessions using ballr

```
library(tidyverse)
nba_pos_stats <-
    read_csv("https://shorturl.at/mFGY2")
# Find rows for players indicating a full season worth of stats
tot_players <- nba_pos_stats %>% filter(tm == "TOT")
# Stack this dataset with players that played on just one team
nba_player_stats <- nba_pos_stats %>%
    filter(!(player %in% tot_players$player)) %>%
    bind_rows(tot_players)
# Filter to only players with at least 125 minutes played
nba_filtered_stats <- nba_player_stats %>% filter(mp >= 125)
head(nba_filtered_stats)
```

```
## # A tibble: 6 × 31
```

##		player	pos	age	tm	g	gs	mp	fg	fga	fgper…¹	х3р	х3ра
##		<chr></chr>	<chr></chr>	<dbl></dbl>	<chr></chr>	<dbl></dbl>							
##	1	Precious …	С	22	TOR	73	28	1725	7.7	17.5	0.439	1.6	4.5
##	2	Steven Ad	С	28	MEM	76	75	1999	5	9.2	0.547	Θ	0
##	3	Bam Adeba	С	24	MIA	56	56	1825	11.1	20	0.557	Θ	0.2
##	4	Santi Ald…	PF	21	MEM	32	Θ	360	7	17.5	0.402	0.8	6.4
##	5	LaMarcus …	С	36	BRK	47	12	1050	11.6	21.1	0.55	0.6	2.1
##	2	Gravean A	50	26	мті	66	61	1005	6 0	15 1	0 110	1 2	10 /

15 / 20

Gaussian Mixture Models with mclust

Use the Mclust function to search over 1 to 9 clusters (*K* = G) and the different covariance constraints (i.e. models)

library(mclust)
nba_mclust <- Mclust(dplyr::select(nba_filtered_stats, x3pa, trb))</pre>

We can use the summary() function to display the selection and resulting table of assignments:

```
summary(nba_mclust)
```

```
_____
## Gaussian finite mixture model fitted by EM algorithm
##
##
## Mclust VVI (diagonal, varying volume and shape) model with 3 components:
##
   log-likelihood n df BIC
##
                                   ICL
        -2459.03 483 14 -5004.581 -5141.138
##
##
## Clustering table:
    1 2 3
##
##
  52 276 155
```

Display the BIC for each model and number of clusters



plot(nba_mclust, what = 'classification')





Number of components

How do the cluster assignments compare to the positions?

We can again compare the clustering assignments with player positions:

table("Clusters" = nba_mclust\$classification, "Positions" = nba_filtered_stats\$pos)

##	Positions											
##	Clusters	С	C-PF	ΡF	PF-SF	PG	PG-SG	SF	SF-SG	SG	SG-PG	SG-SF
##	1	43	Θ	9	Θ	0	0	0	0	0	0	0
##	2	3	Θ	28	Θ	84	0	54	5	96	3	3
##	3	39	2	56	1	8	1	38	0	9	0	1

What about the cluster probabilities?

```
nba_player_probs <- nba_mclust$z</pre>
colnames(nba_player_probs) <-</pre>
  paste0('Cluster ', 1:3)
nba_player_probs <- nba_player_probs %>%
  as_tibble() %>%
  mutate(player =
           nba filtered stats$player) %>%
  pivot_longer(contains("Cluster"),
               names_to = "cluster",
               values_to = "prob")
nba_player_probs %>%
  ggplot(aes(prob)) +
  geom_histogram() +
  theme_bw() +
```

facet_wrap(~ cluster, nrow = 2)



Which players have the highest uncertainty?

```
nba_filtered_stats %>%
 mutate(cluster =
          nba_mclust$classification,
         uncertainty =
           nba_mclust$uncertainty) %>%
 group_by(cluster) %>%
 arrange(desc(uncertainty)) %>%
 slice(1:5) %>%
 ggplot(aes(y = uncertainty,
             x = reorder(player,
                         uncertainty))) +
 geom_point() +
 coord_flip() +
 theme_bw() +
 facet_wrap(~ cluster,
             scales = 'free_y', nrow = 3)
```



20 / 20